

For a binomial distribution if  $n$  the number of trials is very large but the mean  $np$  stays roughly constant (meaning that the probability,  $p$  becomes very small) we have events which are relatively rare, but still occur with some regularity. For example, traffic accidents looked over time at a particular intersection or defects produced in a continuous output stream. The resulting distribution is a special limit of the binomial distribution and is named after the French mathematician, Simeon Poisson. A binomial distribution with  $n$  trials and a probability of success,  $p$ , on any one trial gives the following probability function for  $x$  successes in the  $n$  attempts,

$$P(x) = \binom{n}{x} p^x (1-p)^{n-x}.$$

Taking logarithms,

$$\ln[x!P(x)] = \ln(n!) - \ln[(n-x)!] + x \ln(p) + (n-x) \ln(1-p)$$

Replacing  $p$  by  $p = \frac{\langle x \rangle}{n} = \mu$ , gives

$$\ln[x!P(x)] = \ln(n!) - \ln[(n-x)!] + x \ln(\mu) - x \ln(n) + (n-x) \ln\left(1 - \frac{\mu}{n}\right)$$

Stirling's Formula gives the asymptotic form of  $n!$ ,  $n! \approx \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$ , so that

$$\ln(n!) \approx n \ln(n) - n \ln(e) + \frac{1}{2} \ln(n) + \frac{1}{2} \ln(2\pi) = n \ln(n) - n + \frac{\ln(n)}{2} + \frac{\ln(2\pi)}{2}$$

$$\begin{aligned} \ln[(n-x)!] &\approx (n-x) \ln(n-x) - (n-x) + \frac{\ln(n-x)}{2} + \frac{\ln(2\pi)}{2} \\ &\approx n \left(1 - \frac{x}{n}\right) \ln\left[n \left(1 - \frac{x}{n}\right)\right] - n + x + \frac{\ln\left[n \left(1 - \frac{x}{n}\right)\right]}{2} + \frac{\ln(2\pi)}{2} \\ &\approx n \left(1 - \frac{x}{n}\right) \ln(n) + n \left(1 - \frac{x}{n}\right) \ln\left(1 - \frac{x}{n}\right) - n + x + \frac{\ln(n)}{2} + \frac{\ln\left(1 - \frac{x}{n}\right)}{2} + \frac{\ln(2\pi)}{2} \end{aligned}$$

The Maclaurin series for  $\ln(1+q) = q - \frac{q^2}{2} + \frac{q^3}{3} - \frac{q^4}{4} + \frac{q^5}{5} \dots$  converges for  $|q| < 1$ .

Keeping terms to order  $\frac{1}{n}$ ,

$$\begin{aligned} \ln[(n-x)!] &\approx n \ln(n) - n + \frac{\ln(n)}{2} - x \ln(n) + x - x + \frac{\ln(2\pi)}{2} + \frac{x^2}{n} - \frac{x}{2n}, \text{ so that} \\ &\approx n \ln(n) - n + \frac{\ln(n)}{2} - x \ln(n) + \frac{\ln(2\pi)}{2} + \frac{x^2}{n} - \frac{x}{2n} \end{aligned}$$

$$\begin{aligned} \ln(n!) - \ln[(n-x)!] &\approx n \ln(n) - n + \frac{\ln(n)}{2} + \frac{\ln(2\pi)}{2} - n \ln(n) + n - \frac{\ln(n)}{2} + x \ln(n) - \frac{\ln(2\pi)}{2} - \frac{x^2}{n} + \frac{x}{2n} \\ &\approx x \ln(n) - \frac{x^2}{n} + \frac{x}{2n} \end{aligned}$$

Using this result in the expansion of  $\ln[x!P(x)]$  gives

$$\begin{aligned} \ln[x!P(x)] &\approx x \ln(n) - \frac{x^2}{n} + \frac{x}{2n} + x \ln(\mu) - x \ln(n) + (n-x) \ln\left(1 - \frac{\mu}{n}\right) \\ &\approx x \ln(\mu) + n \left(1 - \frac{x}{n}\right) \ln\left(1 - \frac{\mu}{n}\right) - \frac{x^2}{n} + \frac{x}{2n} \\ &\approx x \ln(\mu) - \mu \left(1 - \frac{x}{n}\right) - \frac{x^2}{n} + \frac{x}{2n} = x \ln(\mu) - \mu + O\left(\frac{1}{n}\right) \\ &\approx \ln(\mu^x) + \ln(e^{-\mu}) + O\left(\frac{1}{n}\right) = \ln(e^{-\mu} \mu^x) + O\left(\frac{1}{n}\right) \end{aligned}$$

So as  $n \rightarrow \infty$ ,  $\ln[x!P(x)] \rightarrow \ln(e^{-\mu} \mu^x)$ , or  $x!P(x) \rightarrow e^{-\mu} \mu^x$ . This yields the Poisson probability distribution.

$$P(x) = \frac{e^{-\mu} \mu^x}{x!} \quad (1)$$

Since  $n \rightarrow \infty$ ,  $x$  can be any whole number. To see that this defines a valid probability distribution, use the Maclaurin series for the exponential function.

$$\sum_{x=0}^{\infty} P(x) = \sum_{x=0}^{\infty} \frac{e^{-\mu} \mu^x}{x!} = e^{-\mu} \sum_{x=0}^{\infty} \frac{\mu^x}{x!} = e^{-\mu} e^{\mu} = 1$$

The mean of the Poisson distribution is calculated as follows:

$$\begin{aligned} \mu_x = \langle x \rangle &= \sum_{x=0}^{\infty} xP(x) = \sum_{x=0}^{\infty} \frac{x e^{-\mu} \mu^x}{x!} = e^{-\mu} \sum_{x=1}^{\infty} \frac{x \mu^x}{x(x-1)!} = e^{-\mu} \sum_{x=1}^{\infty} \frac{\mu^x}{(x-1)!} \\ &= e^{-\mu} \sum_{x=1}^{\infty} \frac{\mu^{x-1} \mu^1}{(x-1)!} = \mu e^{-\mu} \sum_{x=1}^{\infty} \frac{\mu^{x-1}}{(x-1)!} = \mu e^{-\mu} \sum_{x=0}^{\infty} \frac{\mu^x}{x!} = \mu e^{-\mu} e^{\mu} = \mu \end{aligned}$$

Thus, the mean is just  $\mu$ . But this was to be expected! We started with a binomial distribution of mean  $\mu$  and it wasn't supposed to change as  $n$  increased. To calculate the variance we calculate the second moment about the origin.

$$\begin{aligned} \langle x^2 \rangle &= \langle x^2 - x \rangle + \langle x \rangle = \langle x(x-1) \rangle + \mu \\ \langle x(x-1) \rangle &= \sum_{x=0}^{\infty} x(x-1)P(x) = \sum_{x=0}^{\infty} \frac{x(x-1)e^{-\mu} \mu^x}{x!} = e^{-\mu} \sum_{x=2}^{\infty} \frac{x(x-1)\mu^x}{x(x-1)(x-2)!} = e^{-\mu} \sum_{x=2}^{\infty} \frac{\mu^x}{(x-2)!} \\ &= e^{-\mu} \sum_{x=2}^{\infty} \frac{\mu^{x-2} \mu^2}{(x-2)!} = \mu^2 e^{-\mu} \sum_{x=2}^{\infty} \frac{\mu^{x-2}}{(x-2)!} = \mu^2 e^{-\mu} \sum_{x=0}^{\infty} \frac{\mu^x}{x!} = \mu^2 e^{-\mu} e^{\mu} = \mu^2 \end{aligned}$$

So,  $\langle x^2 \rangle = \mu^2 + \mu$ ;  $\sigma_x^2 = \langle x^2 \rangle - \langle x \rangle^2 = \mu^2 + \mu - \mu^2 = \mu$ . The variance of a Poisson distribution is the same as its mean.